### Lower Bounds on the Rate of Convergence for Accept-Reject-Based Markov Chains

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#### Setting

We have a target distribution  $\Pi$  on  $\mathbb{R}^d$  possibly depending on data of size n (e.g. Bayesian posteriors) with Lebesgue density  $\pi$  with support  $\Theta \subseteq \mathbb{R}^d$ .

We want to generate representative samples  $\theta_1, \ldots, \theta_T$  from  $\Pi$  to approximate expectations (e.g. predictions and inference in Bayesian statistics)

$$\frac{1}{T}\sum_{t=1}^{T}g(\theta_t) \approx \int g(\theta)\pi(\theta)d\theta.$$

#### **Bayesian Applications**

Consider the Bayesian model

$$\theta \sim \pi_0$$
  
 $X_1, Y_1, \dots, X_n, Y_n | \theta \sim p_{n,\theta}(\cdot)$ 

Observe  $x_1, y_1, \ldots, x_n, y_n$  and the Bayesian posterior density

$$\pi_n(\theta) \propto \pi_0(\theta) p_{n,\theta}(x_1, y_1, \dots, x_n, y_n).$$

Often the density is intractable and the normalizing constant is difficult to approximate.

Example:

Bayesian GLM's (i.e. logistic regression)

### Difficulties in Sampling

Independent sampling methods are infeasible for modern complex target distributions.

- Inverting the distribution function / transformation methods (e.g Normal distributions)
- Rejection sampling (Gamma distributions, etc)
- Normalizing constant of π is often unknown (difficult for importance sampling)

MCMC approach: Simulate a Markov chain for sufficiently long with stationary distribution  $\Pi$ .

Main concern: How long do we need to simulate the Markov chain?

#### Accept-reject-based Markov chains

Generate  $\theta_t | \theta_0 \sim P^t(\theta_0, \cdot)$  in discrete-time using a proposal  $Q(\cdot, \cdot)$  by

$$\theta_t | \theta_{t-1} = \begin{cases} \theta'_t, & \text{if } u_t \leq a(\theta_{t-1}, \theta'_t) \\ \theta_{t-1}, & \text{else} \end{cases}$$

where  $\theta_t' | \theta_{t-1} \sim Q(\theta_{t-1}, \cdot)$  and  $u_t \sim \text{Unif}(0, 1)$ .

Choose  $a(\cdot, \cdot)$  in a way so  $\Pi$  is invariant (not necessarily reversible).

**Examples:** Metropolis-Hastings [Metropolis et al., 1953, Hastings, 1970, Tierney, 1998], Barker's Barker [1964], non-reversible Metropolis-Hastings Bierkens [2015]

### Metropolis-Hastings

If  $Q(\cdot, \cdot)$  with transition density  $q(\cdot, \cdot),$  Metropolis-Hastings:

$$a(\theta_{t-1}, \theta'_t) = \frac{\pi \left(\theta'_t\right) q \left(\theta'_t, \theta_{t-1}\right)}{\pi \left(\theta_{t-1}\right) q \left(\theta_{t-1}, \theta'_t\right)} \wedge 1$$

Optimal in a Peskun sense [Tierney, 1998].



Figure: Arianna Rosenbluth, Nicholas Metropolis, Keith Hastings, and Luke Tierney

#### Drawbacks

#### Drawbacks:

- Requires choosing an acceptance function  $a(\cdot, \cdot)$ .
- Requires choosing and tuning a proposal  $Q(\cdot, \cdot)$ .
  - Independent proposal
  - Random-walk proposals (RWM):  $\theta'_t \sim N(\theta_{t-1}, hI_d)$
  - Discretize Langevin diffusions (MALA):  $\theta'_t \sim N(\theta_{t-1} + h\nabla \log(\pi(\theta_{t-1}))/2, hI_d)$

Can be unreliable if the proposal is chosen poorly.

#### The Problem

**Problem:** Practitioners often require tuning proposals by trial and error to avoid poor empirical performance in many applications.

We want to contribute to existing tools for choosing tuning parameters:

- Optimal scaling for RWM, MALA Roberts et al. [1997], Roberts and Rosenthal [1998]
- Adaptive algorithms Haario et al. [2001]
- Convergence analysis

### Trial and Error

 Trial and error: Tune the algorithm, run the algorithm, monitor acceptances, and restart if acceptances are low.



#### Drawbacks:

• Computationally intensive and time consuming.

#### Convergence Rate Upper bounds

Recent interest in (geometric) convergence rate upper bounds for MCMC algorithms in terms of the problem size d, n [Belloni and Chernozhukov, 2009, Ekvall and Jones, 2021, Johndrow et al., 2019, Qin and Hobert, 2019, Rajaratnam and Sparks, 2015, Yang et al., 2016].

Also an interest in new coupling techniques in Wasserstein distances since they appear to scale better in high dimensions [Hairer et al., 2014, Qin and Hobert, 2019, 2021].

**Difficulties:** Explicit convergence rates for Metropolis-Hastings in TV or Wasserstein is a challenging problem and the convergence rates are largely unknown [Jarner and Hansen, 2000, Hairer et al., 2014].

Some mixing time bounds and recent work in the area [Dwivedi et al., 2018, Andrieu et al., 2022].

Use lower bounds on the convergence rates to aid practitioners in understanding which tuning parameters may cause the algorithms to *fail* produce a representative sample in an available number of iterations in terms of d, n.

Drawbacks: Does not tell you when the algorithm performs well.

**Related literature:** Exact rates for independence samplers [Wang, 2022, Brown and Jones, 2021]. Necessary conditions for geometric ergodicity [Roberts and Tweedie, 1996].

# Geometric Ergodicity in TV and Wasserstein Distances

#### Intuition: Transportation Distances



Optimally transport all the mass from one probability distribution to the other with cost  $c(\cdot, \cdot)$ .

Examples:  $c(\theta', \omega') = I_{\theta' \neq \omega'}$  and  $c(\theta', \omega') = \|\theta' - \omega'\|$ .

#### Transportation Distances

Let  $\mathcal{C}\left(P^t(\theta,\cdot),\Pi\right)$  be the set of couplings. The Wasserstein distance is defined as

$$\mathcal{W}_{\|\cdot\|}^{p}\left(P^{t}(\theta,\cdot),\Pi\right) = \left(\inf_{\xi\in\mathcal{C}(P^{t}(\theta,\cdot),\Pi)}\int\left\|\theta'-\omega'\right\|^{p}d\xi(\theta',\omega')\right)^{1/p}$$

in comparison to

$$\left\|P^{t}(\theta, \cdot) - \Pi\right\|_{\mathsf{TV}} = \inf_{\xi \in \mathcal{C}(P^{t}(\theta, \cdot), \Pi)} \int I_{\theta' \neq \omega'} d\xi(\theta', \omega')$$



Figure: Leonid Kantorovich, Leonid Vaseršteĭn, Cédric Villani

#### Geometric Ergodicity

An accept-reject-based Markov chain is  $(\rho, M)$ -geometrically ergodic if for  $\rho \in (0, 1)$  and a function  $M(\cdot)$ , we have for every initialization  $\theta \in \Theta$ ,

$$\left\|P^t(\theta,\cdot) - \Pi\right\|_{\mathsf{TV}} \le M(\theta)\rho^t$$

and  $\left( \left\| \cdot \right\|, p, \rho, M \right)$  -geometrically ergodic if

$$\mathcal{W}_{\parallel \cdot \parallel}^p(P^t(\theta, \cdot), \Pi) \le M(\theta)\rho^t.$$

**Motivation:** Upper bounds on convergence rates in Wasserstein distances tend to scale better in large problem sizes [Hairer et al., 2014, Qin and Hobert, 2019].

Geometric Ergodicity Can be Slow to Converge

Convergence can be slow if  $\rho \approx 1$ .

- Generated samples are not trustworthy
- Suggests unreliable estimators from the Markov chain



## Lower Bounds

#### Lower bounds on the TV Convergence Rate

Theorem (Theorem 1, 2 [Brown and Jones, 2022])

For any 
$$heta \in \Theta$$

$$\left\|P^{t}(\theta, \cdot) - \Pi\right\|_{TV} \ge \left[1 - A(\theta)\right]^{t}$$

where  $A(\theta)=\int a(\theta,\theta')Q(\theta,d\theta').$  If  $(\rho,M)$  -geometrically ergodic so

$$\left\|P^t(\theta, \cdot) - \Pi\right\|_{TV} \le M(\theta)\rho^t,$$

then

$$1 - \inf_{\theta \in \Theta} A(\theta) \le \rho$$

- Method independent (e.g. drift and minorization, coupling)
- Does not require reversiblility

#### Comparing Algorithms

If P is  $(\rho, M)$ -geometrically ergodic and  $A(\theta) \leq A_{MH}(\theta)$  (Peskun ordered) where  $A_{MH}$  is the version for Metropolis-Hastings, then

$$1 - \inf_{\theta \in \Theta} A_{MH}(\theta) \le \rho.$$

#### Lower Bounds for Wasserstein Distances

Theorem (Theorem 4, 5 [Brown and Jones, 2022]) If  $\pi$  is bounded, then there is a  $C_{d,\pi} > 0$  so every  $\theta \in \Theta$ 

$$\mathcal{W}_{\parallel,\parallel}^p(P^t(\theta,\cdot),\Pi) \ge C_{d,\pi} \left[1 - A(\theta)\right]^{t\left(1 + \frac{1}{d}\right)}$$

If  $(\left\|\cdot\right\|,p,\rho,M)\text{-geometrically ergodic so}$ 

 $\mathcal{W}_{\|\cdot\|}^p(P^t(\theta,\cdot),\Pi) \le M(\theta)\rho^t,$ 

then

$$1 - \inf_{\theta \in \Theta} A(\theta) \le \rho^{\frac{d}{d+1}}.$$

Similar to total variation in high dimensions

#### The Approach

- Find problematic point: The maximum of the target density  $\theta^*$  can be a problematic for Metropolis-Hastings.
- Study the computational complexity: Study how  $A(\theta^*) \to 0$  with the problem size d, n.
- Use lower bounds: Lower bounds in TV give  $1 \rho \le A(\theta^*) \to 0$  with the problem size d, n.

Focus on M-H and focus on TV since Wasserstein will be similar.

# Applications Under Concentration

#### Application: RWM for Log-concave targets

Consider  $\pi \propto \exp(-f)$  and suppose M-H with RWM proposal  $\theta'_t \sim N(\theta_{t-1}, hI_d)$  is  $(\rho, M)$ -geometrically ergodic.

Corollary (Corollary 1, 2 [Brown and Jones, 2022]) If  $f(\cdot) - \frac{1}{2\xi} \|\cdot\|_2^2$  is convex on  $\mathbb{R}^d$ ,  $1 - \rho \leq \frac{1}{(h/\xi + 1)^{d/2}}$ .

• Need to choose  $h \propto \xi/d$  as  $d \to \infty$  to avoid  $\rho \to 1$ 

Many examples: Bayesian GLM with Gaussian priors

Example: Bayesian Logistic Regression with Zellner's g-prior

Consider the posterior  $\Pi_n$  with i.i.d. data  $(Y_i, X_i)_i$ 

 $Y_i | X_i, \beta \sim \mathsf{Bern}\left(\mathsf{sigmoid}(X_i^T \beta)\right) \quad \beta \sim N_d\left(0, g\left(X^T X\right)^{-1}\right)$ 

Assume  $(X_{i,j})_{i,j}$  are i.i.d. random variables with zero mean, unit variance, and a finite fourth moment.

Suppose M-H is  $(\rho_n, M_n)$  geometrically ergodic each n with a RWM proposal  $\theta'_t \sim N(\theta_{t-1}, hI_d)$ .

Example: Bayesian Logistic Regression with Zellner's g-prior

Proposition (Proposition 5 [Brown and Jones, 2022])

Suppose  $n \to \infty$  with  $d_n/n \to \gamma \in (0,1)$ . Then w.p. 1 and large n,

$$1 - \rho_n \le \frac{1}{\left(\frac{hn(1-\sqrt{\gamma})^2}{2g} + 1\right)^{d_n/2}}.$$

Choose  $h \propto 1/(dn)$  or  $\lim_{n,d_n} \rho_n = 1$  can be rapid!

#### Numerical Simulation

- Generate repeatedly 50 times artificial data with increasing dimensions d, n = 4d:  $(d, n) \in \{(2, 8), ..., (14, 56)\}$ .
- $\blacksquare$  Use optimization and Monte Carlo to estimate  $A(\beta_n^*)$  and lower bounds.

• 
$$h = .6, 2.38^2/d, 1/(dn).$$



#### More General Lower Bounds Under Concentration

Suppose conditions (roughly speaking) on  $\pi_n$ :

- local  $\lambda_0^{-1}$ -strongly convex condition
- strict maximum
- sufficient tail decay

Suppose M-H with general proposal  $\theta'_t \sim N(\mu(\theta_{t-1}), hC)$  (i.e. RWM, MALA, Riemannian manifold MALA [Girolami and Calderhead, 2011]) is  $(\rho_n, M_n)$ -geometrically ergodic for each n.

#### Lower Bounds Under Concentration

Proposition (Proposition 6, 8, [Brown and Jones, 2022]) Under conditions on  $\pi_n$  and  $d_n \leq n^{\kappa}$ ,  $\kappa \in (0, 1)$ , then for large  $(n, d_n)$ ,  $1 - \rho_n \leq \left(\frac{\lambda_0}{nh}\right)^{d_n/2} \frac{2}{\det(C)^{1/2}}.$  (1)

- For a large class of proposals, if h, C do not depend carefully on n, then  $\lim_{(n,d_n)\to\infty} \rho_n = 1$  rapidly!
- Similar bound holds for any bounded proposal.

Application: Flat prior Bayesian logistic regression

Consider  $\pi_n$  for the model with i.i.d. data  $(Y_i,X_i)_i$  with  $\|X_i\|_2 \leq 1$  w.p. 1

$$Y_i | X_i, \beta \sim \text{Bern}\left(\text{sigmoid}(X_i^T \beta)\right) \qquad \beta \propto 1$$

Suppose M-H with RWM proposal  $\theta'_t \sim N(\theta_{t-1}, hI_d)$  is  $(\rho_n, M_n)$ -geometrically ergodic.

Theorem (Theorem 3 [Brown and Jones, 2022])

In fixed dimension d and under conditions so the target exists [Chen and Shao, 2000] and the MLE is consistent w.p. 1 and  $X_i^T u \neq 0$  if  $u \neq 0$ . There is a  $\lambda_0 > 0$  such that w.p. 1, for large n,

$$1 - \rho_n \le 2 \left(\frac{\lambda_0}{nh}\right)^{d/2}$$

Can choose  $h \propto 1/n$  to avoid  $\lim_{n} \rho_n = 1$ .

#### Numerical Simulation

• Generate repeatedly 50 times artificial data with  $(d, n) = \in \{(10, 100), \dots, (10, 400)\}.$ 

• 
$$h = .1, 5/n, 1/n, .1/n$$



# Comparison to Spectral Methods

#### Comparison to Spectral Methods

Proposition (Proposition, Proposition 8 [Brown and Jones, 2022])

If P is reversible and there is a  $\rho \in (0,1)$ , for every  $\mu$  with  $d\mu/d\Pi \in L^2(\Pi)$ , there is a  $M_\mu > 0$  such that

 $\mathcal{W}^{1}_{\|\cdot\|\wedge 1}\left(\mu P^{t},\Pi\right) \leq M_{\mu}\rho^{t}.$ 

If  $A(\cdot)$  is upper semicontinuous, then

$$1 - \inf_{\theta \in \Theta} A(\theta) \le \rho.$$

Based on previous results [Hairer et al., 2014]

#### Summary

#### Choose tuning parameters carefully!

Manuscript is on arXiv (submitted to Annals of Statistics): https://arxiv.org/abs/2212.05955

- Developed similar general lower bounds in both total variation and Wasserstein distances in terms of the acceptance probability.
- Studied applications in Bayesian logistic regression for choosing tuning parameters which scale to large problem sizes to avoid the convergence rate rapidly tending to 1.

#### References I

Christophe Andrieu, Anthony Lee, Sam Power, and Andi Q. Wang. Explicit convergence bounds for metropolis markov chains: isoperimetry, spectral gaps and profiles, 2022.

- A. A. Barker. Monte Carlo calculations of the radial distribution functions for a proton-electron plasma. *Australian Journal of Physics*, 18:119–132, 1964.
- Alexandre Belloni and Victor Chernozhukov. On the computational complexity of MCMC-based estimators in large samples. *The Annals of Statistics*, 37(4):2011–2055, 2009.
- Joris Bierkens. Non-reversible Metropolis-Hastings. *Statistics and Computing*, 26(6):1213–1228, 2015.
- Austin Brown and Galin L. Jones. Exact convergence analysis for Metropolis-Hastings independence samplers in Wasserstein distances. *preprint arXiv:2111.10406*, 2021.

#### References II

- Austin Brown and Galin L. Jones. Lower bounds on the rate of convergence for accept-reject-based markov chains. *preprint arXiv*, 2022.
- Ming-Hui Chen and Q. Shao. Propriety of posterior distribution for dichotomous quantal response models. *Proceedings of the American Mathematical Society*, 129(1):293 302, 2000.
- Raaz Dwivedi, Yuansi Chen, Martin J Wainwright, and Bin Yu.
  Log-concave sampling: Metropolis-Hastings algorithms are fast!
  In Proceedings of the 31st Conference On Learning Theory,
  volume 75 of Proceedings of Machine Learning Research, pages 793–797, 2018.
- Karl Oskar Ekvall and Galin L. Jones. Convergence analysis of a collapsed Gibbs sampler for Bayesian vector autoregressions. *Electronic Journal of Statistics*, 15:691 – 721, 2021.

#### References III

Mark Girolami and Ben Calderhead. Riemann manifold Langevin and Hamiltonian Monte Carlo methods. *Journal of the Royal Statistical Society. Series B (Statistical Methodology)*, 73(2): 123–214, 2011.

- Heikki Haario, Eero Saksman, and Johanna Tamminen. An adaptive Metropolis algorithm. *Bernoulli*, 7(2):223 242, 2001.
- Martin Hairer, Andrew M. Stuart, and Sebastian J. Vollmer. Spectral gaps for a Metropolis–Hastings algorithm in infinite dimensions. *The Annals of Applied Probability*, 24:2455–2490, 2014.
- W. K. Hastings. Monte Carlo sampling methods using Markov chains and their applications. *Biometrika*, 57(1), 1970.
- Søren F. Jarner and Ernst Hansen. Geometric ergodicity of Metropolis algorithms. *Stochastic Processes and their Applications*, 85:341–361, 2000.

#### References IV

- James E. Johndrow, Aaron Smith, Natesh Pillai, and David B. Dunson. MCMC for imbalanced categorical data. *Journal of the American Statistical Association*, 114(527):1394–1403, 2019.
- N. Metropolis, A. W. Rosenbluth, M. N. Rosenbluth, A. H. Teller, and E. Teller. Equations of state calculations by fast computing machine. *Journal of Chemical Physics*, 21, 1953.
- Qian Qin and James P Hobert. Convergence complexity analysis of Albert and Chib's algorithm for Bayesian probit regression. *Annals of Statistics*, 47:2320–2347, 2019.
- Qian Qin and James P. Hobert. On the limitations of single-step drift and minorization in Markov chain convergence analysis. *The Annals of Applied Probability*, 31(4):1633 – 1659, 2021.

#### References V

- Bala Rajaratnam and Doug Sparks. MCMC-based inference in the era of big data: A fundamental analysis of the convergence complexity of high-dimensional chains. *preprint* arXiv:1508.00947, 2015.
- G. O. Roberts, A. Gelman, and W. R. Gilks. Weak convergence and optimal scaling of random walk metropolis algorithms. *The Annals of Applied Probability*, 7(1):110–120, 1997.
- Gareth O. Roberts and Jeffrey S. Rosenthal. Optimal scaling of discrete approximations to langevin diffusions. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)*, 60 (1):255–268, 1998.
- Gareth O. Roberts and Richard L. Tweedie. Geometric convergence and central limit theorems for multidimensional Hastings and Metropolis algorithms. *Biometrika*, 83:95–110, 1996.

#### References VI

- Luke Tierney. A note on Metropolis-Hastings kernels for general state spaces. *The Annals of Applied Probability*, 8:1–9, 1998.
- Guanyang Wang. Exact convergence analysis of the independent Metropolis-Hastings algorithms. *Bernoulli*, 28(3):2012 – 2033, 2022.
- Yun Yang, Martin J. Wainwright, and Michael I. Jordan. On the computational complexity of high-dimensional Bayesian variable selection. *Annals of Statistics*, 44:2497–2532, 2016.